

Multi-bit Decentralized Detection of a Non-cooperative Moving Target Through a Generalized Rao Test

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Abstract—We consider decentralized detection (DD) of an uncooperative moving target via wireless sensor networks (WSNs), measured in zero-mean unimodal noise. To address energy and bandwidth limitations, the sensors use multi-level quantizers. The encoded bits are then reported to a fusion center (FC) via binary symmetric channels. Herein, we propose a generalized Rao (G-Rao) test as a simpler alternative to the generalized likelihood ratio test (GLRT). Further, the asymptotic performance of a trajectory-clairvoyant (multi-bit) Rao test is leveraged to develop an offline and per-sensor quantizer design. Simulations show the appeal of G-Rao test with respect to the GLRT, and the gain in detection by using multiple bits for quantization.

Index Terms—Decentralized detection, Generalized Rao test, GLRT, multibit quantizer, wireless sensor networks.

I. INTRODUCTION

Wireless sensor networks (WSNs) have attracted significant interest due to their adoption in surveillance, environmental monitoring, smart cities, etc. Decentralized detection (DD) is one of the key tasks for a WSN; hence, scientific community has put significant effort in its study in last decades¹ [1]–[5].

Due to energy & bandwidth limitations, sensors are designed to *quantize* their measurements (into one or more bits), before sending them to a fusion center (FC) where a system-wide decision is taken [6], [7]. In this case, the optimal per-sensor design is a local likelihood-ratio (LR) quantization [8], [9], with corresponding selection of sensors’ thresholds being a complex task [1], [2]. Thus the bit(s) sent either embodies the estimated binary event via a sub-optimal rule [10] (in one-bit case) or results from a “dumb” quantization [11].

In both cases sensors’ bits are sent to the FC, where they are fused via an intelligently-designed rule meant to overcome sensors’ limited detection capabilities. Sadly, the target to be detected depends on some unknown parameters. This precludes (global) LR implementation at FC [2], which is then faced to test a composite hypothesis. A commonly-adopted fusion rule in such cases corresponds to the generalized LR test (GLRT) [12]–[14]. Yet, in a case of an unknown static (resp. moving) target with unknown location (resp. trajectory), the GLRT requires a grid search on both the target location (trajectory) and emitted signal domains; this motivates the search for simpler fusion rules.

Hence, recent works have devised a *generalized* Rao test for one-bit DD of uncooperative targets in finite-sample [15] and sequential [16] setups. Sadly, there is useful information lost when only one-bit quantizers are adopted and a notable performance gap can be observed w.r.t. unquantized observations [17]. Accordingly, *multi-level quantization* can be adopted to achieve performance gains at the expenses of a mild complexity increase. Based on this idea, multi-bit DD has been recently considered for the simpler scenario of an unknown signal in Gaussian noise [18], [19], where *multi-bit* GLR (*not in closed-form*) and Rao tests (*in closed-form*) have been devised and an asymptotically-optimal thresholds’ design obtained, via a particle swarm optimization (PSO) [20].

Accordingly, herein we study DD of a non-cooperative moving target via WSNs [11], [12], [18], with sensors using multi-bit quantizers. Our model encompasses: (i) *zero-mean unimodal-symmetric* noise pdfs; (ii) quantized data sent to the FC over error-prone links (emulating energy-limited communications) modeled as binary symmetric channels (BSCs). The resulting test is two-sided with nuisance parameters present only under hypothesis \mathcal{H}_1 , thus making inapplicable the standard Rao test [21]. To circumvent this issue and capitalize multi-level measurements, we devise a *multi-bit form of generalized Rao test* (G-Rao), representing (i) a (computationally-) simpler alternative *fusion rule* to the GLRT and (ii) comprising the one-bit G-Rao devised in [15] as a special case. Also, we propose a quantizer design, based on asymptotic performance maximization of a *trajectory-clairvoyant* (TC) Rao test. Such design is per-sensor, accounts for sensor-FC channel status, and requires neither the target signal nor its trajectory, so it can be computed *offline* via PSO (following [18], [19]). Simulations compare both rules in a practical scenario².

²*Notation* - vectors are denoted with lower-case bold letters, with a_n being the n th element of \mathbf{a} ; finite sets are denoted with upper-case calligraphic letters, e.g. \mathcal{A} ; transpose and expectation are denoted with $(\cdot)^T$ and $\mathbb{E}\{\cdot\}$, respectively; probability mass functions (pmfs) and probability density functions (pdfs) are denoted with $P(\cdot)$ and $p(\cdot)$, respectively, while $P(\cdot|\cdot)$ and $p(\cdot|\cdot)$ their corresponding conditional counterparts; the complementary cumulative distribution function (ccdf) is denoted with $F(\cdot)$; the symbols \sim and $\stackrel{\mathcal{L}}{\sim}$ mean “distributed as” and “asymptotically distributed as”; $\mathcal{N}(\mu, \sigma^2)$ denotes a Gaussian PDF with mean μ and variance σ^2 ; χ_k^2 (resp. $\chi_k^2(\xi)$) denotes a chi-square (resp. a non-central chi-square) pdf with k degrees of freedom (resp. and non-centrality parameter ξ).

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II. PROBLEM STATEMENT

We consider a binary hypothesis test where a collection of sensors $k \in \mathcal{K} \triangleq \{1, \dots, K\}$ are deployed to monitor the absence (\mathcal{H}_0) or presence (\mathcal{H}_1) of a target of interest having a partially-specified spatial signature. In the latter case (i.e. \mathcal{H}_1), the target moves along a fixed direction with constant velocity and continuously radiates an unknown deterministic isotropic signal θ . The emitted signal experiences distance-dependent path-loss and additive noise, before reaching individual sensors. The problem can be summarized as follows:

$$\begin{cases} \mathcal{H}_0 : m_k^t = w_k^t, & k \in \mathcal{K}, \\ \mathcal{H}_1 : m_k^t = \theta g(\mathbf{x}^t, \mathbf{s}_k) + w_k^t; & t = 1, \dots, T \end{cases} \quad (1)$$

In Eq. (1), $m_k^t \in \mathbb{R}$ denotes the k th sensor measurement at instant t and $w_k^t \in \mathbb{R}$ indicates the noise random variable (RV). The RVs w_k^t are assumed (a) statistically independent over space (sensors) and (b) i.i.d. over time. In detail we assume each noise RV has $\mathbb{E}\{w_k^t\} = 0$ and *unimodal symmetric* pdf³ $p_{w_k}(\cdot)$. We underline that reliable estimation of the sensor noise pdf(s) can be achieved based on training data.

By denoting with $\mathbf{x}^0 \in \mathbb{R}^d$ and $\mathbf{v} \in \mathbb{R}^d$ the initial target location and the corresponding velocity, respectively, the target location at time t is given by the parametric expression $\mathbf{x}^t = \mathbf{x}^0 + \mathbf{v}t$. Herein, we make the reasonable assumption that *both* \mathbf{x}^0 and \mathbf{v} are *unknown*. On the other hand, $\mathbf{s}_k \in \mathbb{R}^d$ denotes the *known* k th sensor position, as a result of a sensor self-localization procedure. The pair $(\mathbf{x}^t, \mathbf{s}_k)$ *uniquely* determines the value of $g(\mathbf{x}^t, \mathbf{s}_k)$, here denoting the amplitude attenuation function (AAF)⁴, which models how the signal emitted from the target at t decays as it reaches k th sensor. For instance, when noise RVs are modelled as $w_k^t \sim \mathcal{N}(0, \sigma_{w,k}^2)$, the measurement m_k^t is distributed under \mathcal{H}_0 (resp. \mathcal{H}_1) as $m_k^t | \mathcal{H}_0 \sim \mathcal{N}(0, \sigma_{w,k}^2)$ (resp. $m_k^t | \mathcal{H}_1 \sim \mathcal{N}(\theta g(\mathbf{x}^t, \mathbf{s}_k), \sigma_{w,k}^2)$). For compactness, we define the set of unknowns as $\boldsymbol{\xi} \triangleq \{\theta, \mathbf{x}^0, \mathbf{v}\}$.

By looking at Eq. (1) we observe that the test is *two-sided*, namely $\{\mathcal{H}_0, \mathcal{H}_1\}$ corresponds to $\{\theta = \theta_0, \theta \neq \theta_0\}$ ($\theta_0 = 0$). More important, the *unknown* target position \mathbf{x}^t (equivalently the nuisance parameters $\{\mathbf{x}^0, \mathbf{v}\}$) can be estimated at the FC *only* when $\theta \neq \theta_0$, i.e. when the signal is present (\mathcal{H}_1) [21].

Then, to address bandwidth & energy limited budget in WSNs, we assume that the k th sensor employs a (multi-level) $q(k)$ -bit quantizer⁵, in which the observation m_k^t is compared with a set of quantization thresholds $\{\tau_k(i)\}_{i=0}^{2^{q(k)}}$, determining $2^{q(k)}$ non-intersecting intervals covering the whole \mathbb{R} . Precisely, the corresponding quantizer outcome is mapped into a binary codeword $\mathbf{b}_k^t \in \{0, 1\}^{q(k)}$, where $k = 1, 2, \dots, K$. The quantization intervals are associated to $q(k)$ -bit binary codewords, where $c_t(i) \in \{0, 1\}$ and $\mathbf{c}(i) =$

$[c_1(i) \ \dots \ c_{q(k)}(i)]^T$. Hence, the $q(k)$ -bit quantizer of k th sensor at instant t outputs a codeword defined as:

$$\mathbf{b}_k^t \triangleq \begin{cases} \mathbf{c}(1) & -\infty < m_k^t < \tau_k(1) \\ \mathbf{c}(2) & \tau_k(1) \leq m_k^t < \tau_k(2) \\ \vdots & \vdots \\ \mathbf{c}(2^{q(k)}) & \tau_k(2^{q(k)} - 1) \leq m_k^t < +\infty \end{cases} \quad (2)$$

The codeword of k th sensor is then reported to the FC via an error-prone channel link. The communication process of *each bit* is represented by an independent BSC with (known) bit-error probability (BEP) $P_{e,k}$. A *distorted* codeword \mathbf{y}_k^t will be then received by the FC from k th sensor at time t , whose conditional probability obeys $P(\mathbf{y}_k^t = \mathbf{c}_k(i) | \mathbf{b}_k^t = \mathbf{c}_k(j)) = G_{q(k)}(P_{e,k}, d_{i,j})$, where

$$G_{q(k)}(P_{e,k}, d_{i,j}) \triangleq P_{e,k}^{d_{i,j}} (1 - P_{e,k})^{(q(k) - d_{i,j})}, \quad (3)$$

and $d_{i,j} \triangleq d(\mathbf{c}_k(i), \mathbf{c}_k(j))$ denotes the Hamming distance between codewords $\mathbf{c}_k(i)$ and $\mathbf{c}_k(j)$. For notation compactness, we collect the noisy codewords (the soft-quantized measurements) received from the sensors at time t in the set $\mathcal{Y}^t \triangleq \{\mathbf{y}_1^t \ \dots \ \mathbf{y}_K^t\}$ (recall that $\mathbf{y}_k^t \in \{0, 1\}^{q(k)}$ since codeword lengths may differ among sensors) and all the noisy codewords received from the WSN as $\mathcal{Y}^{1:T} \triangleq \{\mathcal{Y}^1, \dots, \mathcal{Y}^T\}$. Accordingly, the pmf of all the observations, as a function of $\boldsymbol{\xi} = \{\theta, \mathbf{x}^0, \mathbf{v}\}$ is given by $p(\mathcal{Y}^{1:T}; \boldsymbol{\xi}) = \prod_{t=1}^T \prod_{k=1}^K P(\mathbf{y}_k^t; \boldsymbol{\xi})$. The corresponding pmf of the contribution from k th sensor at time t can be expanded as

$$P(\mathbf{y}_k^t; \boldsymbol{\xi}) = \sum_{i=1}^{2^{q(k)}} P(\mathbf{y}_k^t | \mathbf{b}_k^t = \mathbf{c}(i)) P(\mathbf{b}_k^t = \mathbf{c}(i); \boldsymbol{\xi}) \quad (4)$$

The quantizer law reported in Eq. (2) implies the following pmf expression for $P(\mathbf{b}_k^t = \mathbf{c}(i); \boldsymbol{\xi})$

$$P(\mathbf{b}_k^t = \mathbf{c}(i); \boldsymbol{\xi}) = \Pr\{\tau_k(i-1) \leq m_k^t < \tau_k(i)\} = F_{w_k}(\tau_k(i-1) - \theta g(\mathbf{x}^t, \mathbf{s}_k)) - F_{w_k}(\tau_k(i) - \theta g(\mathbf{x}^t, \mathbf{s}_k)) \quad (5)$$

where $F_{w_k}(\cdot)$ denotes the cdf of w_k^t .

After receiving $\mathcal{Y}^{1:T}$, the FC takes a global decision. The aim of this paper is the derivation of a (computationally) simple test (based on the statistic $\Lambda(\mathcal{Y}^{1:T})$) deciding for \mathcal{H}_1 (resp. \mathcal{H}_0) when the statistic is above (resp. below) the threshold γ , and the design of quantizers for the *whole* WSN.

III. FUSION RULES DESIGN

A widespread approach to handle composite hypothesis testing (viz. accounting for the presence of unknown parameters) resorts to the GLR [22]. For the DD problem at hand, the corresponding decision statistic is obtained by replacing the unknown parameters $\{\theta, \mathbf{x}^0, \mathbf{v}\}$ with their ML estimates $\{\hat{\theta}, \hat{\mathbf{x}}^0, \hat{\mathbf{v}}\}$ (under \mathcal{H}_1) in the log-LR, i.e.

$$\Lambda_G \triangleq \ln \frac{p(\mathcal{Y}^{1:T}; \hat{\theta}, \hat{\mathbf{x}}^0, \hat{\mathbf{v}})}{p(\mathcal{Y}^{1:T}; \theta_0)} = \sum_{t=1}^T \sum_{k=1}^K \ln \frac{P(\mathbf{y}_k^t; \hat{\theta}, \hat{\mathbf{x}}^0, \hat{\mathbf{v}})}{P(\mathbf{y}_k^t; \theta_0)} \quad (6)$$

³Noteworthy examples of such pdfs are the Gaussian, Laplace, Cauchy and generalized Gaussian distributions with zero mean [22].

⁴We underline that our results apply to any suitably-defined AAF describing the spatial signature of the target to be detected.

⁵Herein, for simplicity, we focus on deterministic quantizers, leaving the more general case of probabilistic quantizers [23] to future studies.

where $\theta_0 = 0$ and $\{\hat{\theta}, \hat{\mathbf{x}}^0, \hat{\mathbf{v}}\}$ are the maximum-likelihood estimates (MLEs), namely $\{\hat{\theta}, \hat{\mathbf{x}}^0, \hat{\mathbf{v}}\} \triangleq \arg \max_{\theta, \mathbf{x}^0, \mathbf{v}} p(\mathcal{Y}^{1:T}; \theta, \mathbf{x}^0, \mathbf{v})$. Note that the MLEs of such unknown parameters do not possess a closed-form. As a result, searching for the solution $\{\hat{\theta}, \hat{\mathbf{x}}^0, \hat{\mathbf{v}}\}$ may require a *huge computational burden*.

Hence, inspired by the approach in [15], [16], the G-Rao test statistic is considered here. The typical Rao test is known to be asymptotically equivalent to the GLRT *in weak-signal condition*⁶, but with a lower computational complexity than the latter one. Referring to Eq. (1), if $\{\mathbf{x}_0, \mathbf{v}\}$ were known, the standard Rao test could be readily computed [22]. Sadly, the above terms are *unavailable* in our test. Still, leveraging Davies approach [21], a family of Rao test statistics can be calculated for different values of $(\mathbf{x}_0, \mathbf{v})$. Then by maximizing such family of statistics (following a ‘‘GLRT-like’’ approach), the G-Rao test statistic can be expressed as

$$\Lambda_R \triangleq \max_{\{\mathbf{x}^0, \mathbf{v}\}} \left(\frac{\partial \ln [p(\mathcal{Y}^{1:T}; \boldsymbol{\xi})]}{\partial \theta} \Big|_{\theta=\theta_0} \right)^2 / I(\theta_0, \mathbf{x}^0, \mathbf{v}), \quad (7)$$

where $I(\theta_0, \mathbf{x}^0, \mathbf{v})$ denotes the *Fisher information* (FI) when $(\mathbf{x}_0, \mathbf{v})$ are known, i.e. $I(\theta, \mathbf{x}^0, \mathbf{v}) \triangleq \mathbb{E} \left\{ \left(\frac{\partial \ln [p(\mathcal{Y}^{1:T}; \theta, \mathbf{x}^0, \mathbf{v})]}{\partial \theta} \right)^2 \right\}$, and evaluated at θ_0 . Henceforth, we briefly describe the necessary steps for obtaining the G-Rao test in explicit form. First, we express the numerator term of Eq. (7) (before evaluating it at $\theta = \theta_0$) as

$$\left(\frac{\partial \ln [p(\mathcal{Y}^{1:T}; \boldsymbol{\xi})]}{\partial \theta} \right)^2 = \left(\frac{\sum_{t=1}^T \sum_{k=1}^K g(\mathbf{x}^t, \mathbf{s}_k) \frac{\sum_{i=1}^{2^{q(k)}} P(\mathbf{y}_k^t | \mathbf{b}_k^t = \mathbf{c}(i)) \rho(\mathbf{b}_k^t = \mathbf{c}(i); \boldsymbol{\xi})}{\sum_{i=1}^{2^{q(k)}} P(\mathbf{y}_k^t | \mathbf{b}_k^t = \mathbf{c}(i)) P(\mathbf{b}_k^t = \mathbf{c}(i); \boldsymbol{\xi})}}{\sum_{i=1}^{2^{q(k)}} P(\mathbf{y}_k^t | \mathbf{b}_k^t = \mathbf{c}(i)) P(\mathbf{b}_k^t = \mathbf{c}(i); \boldsymbol{\xi})}} \right)^2 \quad (8)$$

where the auxiliary definition $\rho(\mathbf{b}_k^t = \mathbf{c}(i); \boldsymbol{\xi}) \triangleq p_{w_k}(\tau_k(i-1) - \theta g(\mathbf{x}^t, \mathbf{s}_k)) - p_{w_k}(\tau_k(i) - \theta g(\mathbf{x}^t, \mathbf{s}_k))$ has been employed. Secondly, we obtain the explicit form of the FI leveraging the result for multi-bit quantized measurements in [19] when replacing h_k with $g(\mathbf{x}^t, \mathbf{x}_k)$. This leads to $I(\theta, \mathbf{x}^0, \mathbf{v}) = \sum_{k=1}^K i_k(\theta, \mathbf{x}^0, \mathbf{v})$, where:

$$i_k(\theta, \mathbf{x}^0, \mathbf{v}) \triangleq \sum_{t=1}^T g^2(\mathbf{x}^t, \mathbf{s}_k) \times \sum_{i=1}^{2^{q(k)}} \frac{\left\{ \sum_{j=1}^{2^{q(k)}} G_{q(k)}(P_{e,k}, d_{i,j}) \rho(\mathbf{b}_k^t = \mathbf{c}(j); \boldsymbol{\xi}) \right\}^2}{\sum_{j=1}^{2^{q(k)}} G_{q(k)}(P_{e,k}, d_{i,j}) P(\mathbf{b}_k^t = \mathbf{c}(j); \boldsymbol{\xi})} \quad (9)$$

⁶That is $|\theta_1 - \theta_0| = c/\sqrt{K}$ for some constant $c > 0$.

Last, we obtain the closed-form Λ_R by using both Eqs. (8)-(9):

$$\Lambda_R = \max_{(\mathbf{x}^0, \mathbf{v})} \frac{1}{I(\theta_0, \mathbf{x}^0, \mathbf{v})} \times \left(\frac{\sum_{t=1}^T \sum_{k=1}^K g(\mathbf{x}^t, \mathbf{s}_k) \frac{\sum_{i=1}^{2^{q(k)}} P(\mathbf{y}_k^t | \mathbf{b}_k^t = \mathbf{c}(i)) \rho(\mathbf{b}_k^t = \mathbf{c}(i); \theta_0)}{\sum_{i=1}^{2^{q(k)}} P(\mathbf{y}_k^t | \mathbf{b}_k^t = \mathbf{c}(i)) P(\mathbf{b}_k^t = \mathbf{c}(i); \theta_0)}}{\sum_{i=1}^{2^{q(k)}} P(\mathbf{y}_k^t | \mathbf{b}_k^t = \mathbf{c}(i)) P(\mathbf{b}_k^t = \mathbf{c}(i); \theta_0)} \right)^2 \quad (10)$$

Despite the seemingly evaluation difficulty, Λ_R can be more easily evaluated than Eq. (6), since G-Rao only requires a grid search on the initial location \mathbf{x}^0 and velocity \mathbf{v} (no need for estimating θ). Precisely, the *complexity* involved is $\mathcal{O}(N_{\mathbf{x}^0} N_{\mathbf{v}} T \sum_{k=1}^K 2^{q(k)})$ based on a 2-D grid, where $N_{\mathbf{x}^0}$ (resp. $N_{\mathbf{v}}$) is the number of initial position (resp. velocity) bins used. Differently, $\mathcal{O}(N_{\theta} N_{\mathbf{x}^0} N_{\mathbf{v}} T \sum_{k=1}^K 2^{q(k)})$ is required for a 3-D grid-based GLR, with an N_{θ} -fold saving for G-Rao.

IV. QUANTIZER DESIGN

The same quantizer design as [19] is precluded in our case, as no (asymptotically-)optimal performance expressions of tests based on the Davies approach are known in the literature [21]. Still, we modify the rationale in [19] to come up with a reasonable and feasible design. With this aim, we first consider the TC Rao statistic $\bar{\Lambda}_R$, which has the knowledge of $(\mathbf{x}^0, \mathbf{v})$ and obeys the following asymptotic⁷ pdf [19]:

$$\bar{\Lambda}_R \underset{a}{\sim} \begin{cases} \chi_1^2 & \text{under } \mathcal{H}_0 \\ \chi_1^2(\lambda_q(\mathbf{x}^0, \mathbf{v})) & \text{under } \mathcal{H}_1 \end{cases} \quad (11)$$

The above result also holds for the TC GLR. The term $\lambda_q(\mathbf{x}^0, \mathbf{v}) \triangleq \theta_1^2 I(\theta_0, \mathbf{x}^0, \mathbf{v})$, (θ_1 is the true value when \mathcal{H}_1 holds) represents the TC non-centrality parameter vs. $(\mathbf{x}^0, \mathbf{v})$ (the larger, the better the performance of TC GLR/Rao test when the known target trajectory is described by $(\mathbf{x}^0, \mathbf{v})$).

Evidently, the non-centrality parameter λ_q *grows monotonically* with the FI evaluated at θ_0 . Such term depends on K quantization threshold vectors (one per sensor), with k th vector being $\boldsymbol{\tau}_k \triangleq [\tau_k(1), \dots, \tau_k(2^{q(k)} - 1)]$, where the two extreme thresholds are fixed as $\tau_k(0) = -\infty$ and $\tau_k(2^{q(k)}) = +\infty$, respectively. Accordingly, we aim to optimize detection performance of the G-Rao detector (as well as GLRT) by solving the following maximization w.r.t. the $\boldsymbol{\tau}_k$'s:

$$\{\boldsymbol{\tau}_k^*\}_{k=1}^K \triangleq \arg \max_{\{\boldsymbol{\tau}_k\}_{k=1}^K} I(\theta_0, \mathbf{x}^0, \mathbf{v}, \{\boldsymbol{\tau}_k\}_{k=1}^K) \quad (12)$$

where the dependence of the FI on the $\boldsymbol{\tau}_k$'s has been highlighted, with a slight abuse of notation. In general, this (TC) procedure would lead to impractical $\boldsymbol{\tau}_k^*$'s, namely depending on $(\mathbf{x}^0, \mathbf{v})$. Still, for our task, the objective admits a *decoupling* into a set of K separate threshold design problems. What's more, each problem is also *independent* of $(\mathbf{x}^0, \mathbf{v})$, namely:

$$\boldsymbol{\tau}_k^* \triangleq \arg \max_{\boldsymbol{\tau}_k} i_k(\theta_0; \boldsymbol{\tau}_k), \quad k = 1, \dots, K \quad (13)$$

⁷Herein, the term ‘‘asymptotic’’ refers both to large-WSN and weak-signal conditions.

We highlight that each maximization is subject to the ordered constraints $\tau_k(i) < \tau_k(i+1)$, for $i = 1, \dots, 2^{q(k)} - 1$. Remarkably, the problem in (13) has the same form as [18], [19], resulting from optimization of Rao (GLR) test performance in a *simpler* unknown signal scenario. Thus, we adopt the method adopted therein, i.e. the PSO, to search each τ_k^* via (13), due to its appeal with high-dimensional and non-convex spaces⁸.

Briefly, PSO is an optimization method [20] resorting to a swarm of $m = 1, 2, \dots, M$ particles to explore⁹ the $(2^q - 1)$ -dimensional space (constrained in each dimension as $\tau(i) \in \{-\tau_{max}, \tau_{max}\}$) in search of a global maximum for each problem in Eq. (13). At ℓ th iteration, the m th particle is characterized by its position τ_m^ℓ (i.e. the objective argument) and velocity vectors ν_m^ℓ (i.e. the improvement direction). The PSO state is summarized by the swarm best position (*sbest* ^{ℓ}), as well as the best individual position achieved by m th particle so far (*pbest* ^{ℓ}). At $(\ell + 1)$ th step, both terms contribute to update each particle velocity $\nu_m^{\ell+1}$ and, indirectly, its position (via $\tau_m^{\ell+1} = \tau_m^\ell + \nu_m^{\ell+1}$). PSO ends when all the particles' velocities (in magnitude) are less than a preset value $\leq \nu_{tol}$.

V. NUMERICAL RESULTS

Herein, we compare multi-bit G-Rao test and GLRT performance in terms of system false alarm ($P_{F_0} \triangleq \Pr\{\Lambda > \gamma | \mathcal{H}_0\}$) and detection probabilities ($P_{D_0} \triangleq \Pr\{\Lambda > \gamma | \mathcal{H}_1\}$). To this end, we consider a 2-D space ($\mathbf{x}_T \in \mathbb{R}^2$) where a WSN of size $K = 9$ is deployed to reveal the presence of an unknown moving target with its initial position located in the (square) surveillance area $\mathcal{L} \triangleq [0, 1]^2$ and moving with velocity within $\mathcal{V} \triangleq [-0.1, 0.1]^2$. W.l.o.g., sensors are displaced in a regular grid covering \mathcal{L} . Concerning the sensing model, we assume $w_k \sim \mathcal{N}(0, \sigma_w^2)$, $k \in \mathcal{K}$, with $\sigma_w^2 = 1$. Also, we consider a power-law AAF $g(\mathbf{x}_T, \mathbf{s}_k) \triangleq 1/\sqrt{1 + (\|\mathbf{x}_T - \mathbf{s}_k\|/0.2)^4}$. The target signal-to-noise ratio (SNR) is defined as $\text{SNR} \triangleq \theta^2/\sigma_w^2$. Herein, we set the *true* values as $\theta = 0.5$ (SNR = -3 dB), $\mathbf{x}_0 = [0, 0.5]^T$, $\mathbf{v} = [0.02, 0.013]^T$ and $T = 20$ in the simulations to gain insight into detectors' performance.

Following Sec. III, Λ_R and Λ_G are evaluated by means of grids for \mathbf{x}_0 , \mathbf{v} and θ . Precisely, \mathbf{x}_0 and \mathbf{v} are searched with $N_{\mathbf{x}_0} = N_{\mathbf{v}} = 100$ grid points uniformly sampling \mathcal{L} and \mathcal{V} , respectively. Differently, the search space of θ (the target signal) is assumed to be $S_\theta \triangleq [-\bar{\theta}, \bar{\theta}]$ ($\bar{\theta} > 0$). The grid points¹⁰ are then chosen as $[-\mathbf{g}_\theta^T \ 0 \ \mathbf{g}_\theta^T]^T$, where \mathbf{g}_θ collects target strengths corresponding to SNR = -10 : 1 : 10 dB.

In Fig. 1 we show P_{D_0} vs. P_{F_0} of both G-Rao and GLRT detectors for (a) $P_{e,k} = P_e = 0$ and (b) $P_{e,k} = P_e = 0.1$, respectively (based on 10^5 Monte Carlo trials). For each case, we report the performance with $q_k = q \in \{1, 2, 3\}$ quantization bits, where thresholds are selected following the rationale elaborated in Sec. IV (via PSO, with parameters $M = 100$, $\tau_{max} = 5$ and $\nu_{tol} = 10^{-6}$). First, it is shown

⁸We remark that other approaches may be pursued, e.g. simulated annealing.

⁹Hereinafter, aiming at improved notation readability, we will omit the sensor index subscript “ k ”.

¹⁰This grid implies $N_\theta = 43$, hence a $43\times$ complexity saving is achieved by G-Rao w.r.t. GLR.

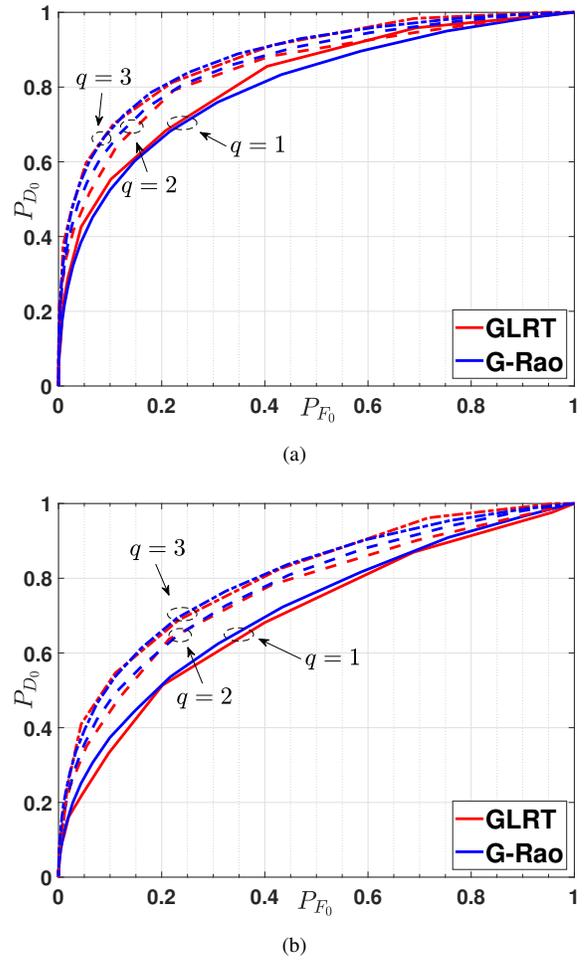


Fig. 1. P_{D_0} vs. P_{F_0} for GLRT and G-Rao; WSN with $K = 9$ sensors, $w_k \sim \mathcal{N}(0, 1)$, SNR = -3 dB, and (a): $P_e = 0$; (b): $P_e = 0.1$.

that the performance of the proposed G-Rao test is practically the same as the GLRT, while having the advantage of a lower computational burden with respect to the latter. Secondly, both fusion rules enjoy a higher detection probability (than the one-bit case) when using multi-bit quantizers.

Still, the presence of channel errors (in our example $P_e = 0.1$) leads to a significant performance loss of both detectors, highlighting the need for either a higher number of sensors (K) or a longer observation interval (T). A deeper numerical analysis will be brought out in the journal version of this paper.

VI. CONCLUDING REMARKS

We devised a G-Rao test for multi-bit DD of a non-cooperative moving target in WSNs. The considered model encompasses unimodal zero-mean symmetric noise, and non-identical BSCs. Our proposal constitutes a simpler alternative to the GLRT, while providing the same performance gains achieved via multi-bit quantization (over a one-bit counterpart). WSN performance was further optimized via the design of PSO-based quantizers, maximizing the asymptotic detection rate of TC Rao rule. Future avenues include the design of G-Rao test for sequential [16] and censoring [24] setups.

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